

# Word Error Rate Measurement in Analog-to-Digital Converters: some statistical considerations about IEEE Std 1241-2000, Annex A

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**Abstract** - The word error rate (WER) in an Analog to Digital Converter (ADC) is the probability of receiving an erroneous code for an input, after correction is made for gain, offset, and linearity errors, and a specified allowance is made for noise. Typical causes of word errors are metastability and timing jitter of comparators within the ADC [1]. The aim of this paper is to give a contribution for a new draft of the IEEE Std 1241. Our attention is directed towards the word error rate estimation and to the Annex A of this standard. New statistical techniques which can better integrate what is sustained in the IEEE standard and in [2] have been proposed. In particular, Student and chi-square distributions have been introduced for a more accurate measurement of the word error rate in the case of  $n$  successive observations.

## I. WER measurement: comments about error (IEEE std 1241 - ANNEX A)

In the Annex A of [1], the presence of statistical errors associated with word-error-rate measurements is discussed. In this context the source of word errors is assumed purely random and both the number of observed word errors and the total number  $n$  of trial samples as statistically significant numbers are considered.

In this case the best estimate of the errors associated with the error rate is given by equation A.1 in [1], through a confidence level that, hypothesizing a Poisson distribution, can be quantified also by:

$$P\{X \leq x\} = \sum_{i=0}^x \frac{\lambda^i e^{-\lambda}}{i!} \quad i = 0, 1, 2, \dots$$

where  $\lambda$  is the mean number of erroneous words with  $n$  trial samples.

In some measurements, where word error rates are very low, other equations are suggested (eq. A3 and A4 in [1]) in the standard. In this case also we suggested this other expression:

$$P\{X = x\} = p(1-p)^{1-i} \quad i = 1, 2, \dots$$

where  $p$  is the probability to have an erroneous word after  $n$  trial samples, and it is given by the probability mass function of the geometrical model [3].

## II. Introduction to the $t$ (and chi-square) distribution for a more accurate evaluation of the measure of the word error rate

### II.A. Recall of theoretical of basis [3], [4]

It is well known that a generic normally distributed measure  $M=N(m,u)$  with expected value  $m$  and standard uncertainty  $u$ , can be expressed in a reduced form using the  $\frac{M-m}{u} = N(0,1)$  normally distributed, with expected value zero and unitary standard uncertainty. It is also known that in the presence of a number  $\nu$  of normal random variables given in the reduced form  $N_1(0,1), \dots, N_\nu(0,1)$ , mutually independent and independent from  $M$ , the sum of squares of such variables, that is

$\chi_\nu^2 = \sum_{i=1}^{\nu} N_i^2(0,1)$  is distributed like a chi-square distribution with  $\nu$  degrees of freedom. The ratio

between the original measure expressed in the reduced form  $\frac{M-m}{u} = N(0,1)$  and the positive square root of  $\chi^2_\nu$  divided by degrees of freedom  $\nu$ , can be expressed by the following variable:

$$T_\nu = \frac{N(0,1)}{\sqrt{\sum_{i=1}^{\nu} N_i^2(0,1)/\nu}} , \text{ with } E\{N_i(0,1), N_j(0,1)\} = 0 \quad i \neq j = 0,1,\dots,\nu \quad (\text{condition of independence}) \quad (1)$$

$T_\nu$  follows a Student distribution (or  $t$  distribution) with  $\nu$  degrees of freedom. Its probability density function can be expressed by

$$f_\nu(t) = \frac{1}{\sqrt{\pi\nu}} \frac{\Gamma[(\nu+1)/2]}{\Gamma[\nu/2]} \frac{1}{(1+t^2/\nu)^{(\nu+1)/2}} \quad -\infty < t < +\infty; \quad \nu = 1,2,\dots \quad (2)$$

where  $\Gamma(\alpha) = \int_0^{+\infty} e^{-x} x^{\alpha-1} dx$  represents the generic gamma function.

In the case of  $\nu = 1$ , Eq. (2) becomes the probability density function of the so-called Cauchy distribution for which neither the expected value nor the variance are defined. It can be also demonstrated that:

$$\lim_{\nu \rightarrow \infty} f_\nu(t) = \frac{1}{\sqrt{2\pi}} e^{-t^2/2}; \quad \lim_{\nu \rightarrow \infty} \text{Var}\{T_\nu\} = \lim_{\nu \rightarrow \infty} \frac{\nu}{\nu-2} = 1 \quad (3)$$

that is, when the number of degrees of freedom tends to infinity, the Student variable tends to the normal reduced variable. The confidence level  $p$  and the corresponding uncertainty interval  $[-t_p, +t_p]$ , centred around zero for the Student variable with  $\nu$  degrees of freedom, is determined as follows:

$$p = P\{-t_p \leq T_\nu \leq t_p\} = \int_{-t_p}^{t_p} f_\nu(t) dt \quad (4)$$

The Student table of quantiles (see for instance GUM [5] – Table 2.2 p.66) offers different values of  $t_p$  for successive numbers of degrees of freedom corresponding to the various confidence levels  $p$ .

## II.B. Estimation of the Word Error Rate for $n$ successive observations

We can consider  $n$  independent successive observations  $(o_1, \dots, o_n)$  of the same measurand (trail samples), obtained by the same measurement process implemented in identical conditions of repeatability. As known we are, in this situation, in the presence of intrinsic random effects on the measurement process (inherent variability of the measurement process). If we hypothesize that each observation of the word error rate is a normally random variable with expected value  $m_o$  and standard uncertainty  $u_o$  then  $O_i = N(m_o, u_o) \quad \forall i = 1, \dots, n$ .

The arithmetic mean  $\bar{O} = \sum_{i=1}^n \frac{o_i}{n} = N\left(m_o, \frac{u_o}{\sqrt{n}}\right)$  is an estimator, that is a random variable with the same expected value  $m_o$  and variance reduced by a factor  $(1/n)$ .

We now introduce the normal arithmetic mean expressed in its reduced form, that is:  $\frac{\bar{O} - m_o}{\frac{u_o}{\sqrt{n}}} = N(0,1)$ . An estimator  $S(\bar{O})$  of  $\frac{u_o}{\sqrt{n}}$  exists, and it is given by the so-called experimental

standard deviation of the mean, according to the following formula:

$$s(\bar{O}) = \sqrt{\frac{\sum_{i=1}^n (o_i - \bar{O})^2}{n(n-1)}} = \frac{u_o}{\sqrt{n}} \sqrt{\frac{\chi^2}{n-1}} \quad (5)$$

being:

$$\chi_{n-1}^2 = \sum_{i=1}^n \frac{(o_i - \bar{O})^2}{u_o^2} = \sum_{i=1}^n N_i^2(0,1) \quad (6)$$

Eq. (6) denotes the well known chi-square with  $(n-1)$  degrees of freedom, with the  $N_i(0,1)$  mutually independent and independent from  $\bar{O}$  and therefore from  $N_o(0,1)$ .

Consequently, it can be verified that:

$$T_{n-1} = \frac{\bar{O} - m_o}{s(\bar{O})} = \frac{(\bar{O} - m_o)/(u_o/\sqrt{n})}{\sqrt{\frac{\sum_{i=1}^n (o_i - \bar{O})^2}{u_o^2}}} = \frac{N_o(0,1)}{\sqrt{\chi_{n-1}^2/(n-1)}} \quad (7)$$

is a Student variable with  $(n-1)$  degrees of freedom.

If now we consider the uncertainty interval introduced in (5), assuming  $\nu = n-1$  and adopting (7), we can write:

$$p = P\{-t_p \leq T_{n-1} \leq t_p\} = P\{m_o - t_p S(\bar{O}) \leq \bar{O} \leq m_o + t_p S(\bar{O})\} \quad (8)$$

which represents an estimation of the confidence level shown in Eq. (4).

The arithmetic mean  $\bar{O}$  can be interpreted as a final measure but in Eq. (8) any useful information about the uncertainty interval are not deduced being  $m_o$  generally unknown and the estimator  $S(\bar{O})$  of Eq. (6) a random variable.

An approximate solution can be obtained by introducing (see GUM [5] G.3.1) the Student variable:

$$T_{n-1} = \frac{M - \bar{o}}{s_o^-}$$

where the observed data  $[o_1, \dots, o_n]$  are introduced. In this case  $\bar{o}$  represents the estimate of  $m_o$  while

$s_o^- = \sqrt{\frac{\sum_{i=1}^n (o_i - \bar{o})^2}{n(n-1)}}$  is the estimate of  $\frac{u_o}{\sqrt{n}}$ , a value of the estimator  $S_o^-$  obtained in the a

particular moment of measurement.

The uncertainty interval of the final measure  $m$  of the word error rate, although approximated because of the adopted estimate procedures, becomes:

$$p = P\{\bar{o} - t_p s_o^- \leq M \leq \bar{o} + t_p s_o^-\} \quad (9)$$

where  $t_p$  is still given by the above recalled Student table.

### II.C. Estimation of the word error rate for the sum of n successive observation

In this case, the adopted measurement model is that of a final measure  $M$  obtained as a result of the sum of  $r$  arithmetic means of observations  $(\bar{o}_1, \bar{o}_2, \dots, \bar{o}_r)$  where the generic  $\bar{O}_i$  is the mean of  $n_i$  observations leading back to random effects (GUM [5] type A) and  $s$  random variables  $M_1, \dots, M_j$  affected by systematic effects (GUM [5] type B):

$$M = \sum_{i=1}^r \bar{O}_i + \sum_{j=1}^s M_j \quad (10)$$

all the variables belonging to the second part of Eq. (10) are hypothesized as independent. A number of degrees of freedom equal to  $\nu_i = n_i - 1$ , is given to each  $\bar{O}_i$ .

Interesting statistical parameters are:

$$\begin{cases} E\{\bar{O}_i\} = m_{oi}, \sqrt{\text{Var}\{\bar{O}_i\}} = u_{oi}/\sqrt{n_i} & \forall i = 1, \dots, r \\ E\{M_j\} = m_j, \sqrt{\text{Var}\{M_j\}} = u_j & \forall j = 1, \dots, s \end{cases}$$

consequently we have:

$$\begin{cases} E\{M\} = \sum_{i=1}^r m_{oi} + \sum_{j=1}^s m_j \\ \sqrt{\text{Var}\{M\}} = u_M = \left[ \sum_{i=1}^r \frac{u_{oi}^2}{n_i} + \sum_{j=1}^s u_j^2 \right]^{1/2} \end{cases} \quad (11)$$

In this case, the disposal data are:

- i) the observed values in the estimation of  $(m_{o1}, m_{o2}, \dots, m_{or})$
- ii) the computed values  $(s_{o1}^-, s_{o2}^-, \dots, s_{or}^-)$ , being  $s_{oi}^- = \sqrt{\frac{\sum_{k=1}^{n_i} (o_{ik} - \bar{o}_i)^2}{n_i(n_i - 1)}}$  the generic estimation of  $\frac{u_{oi}}{\sqrt{n_i}}$ .
- iii) the parameters  $m_j$  and  $u_j$  deduced by a previous type B estimation.

For example, if  $M_j$  is included in the  $[a, b]$  interval, we can hypothesize an uniform distribution with  $m_j = \frac{a+b}{2}$  and  $u_j = \frac{b-a}{2\sqrt{3}}$ . If this data is available, Eq. (11) can be approximated as follows:

$$\begin{cases} E\{M\} = \sum_{i=1}^r o_i + \sum_{j=1}^s m_j \\ \tilde{u}_M = \left[ \sum_{i=1}^r s_{oi}^2 + \sum_{j=1}^s u_j^2 \right]^{1/2} \end{cases} \quad (12)$$

Going back to Eq. (10), if we admit that it is possible to declare, even recalling the central limit theorem, that  $M$  follows a normal distribution, we obtain:

$$M = N(\tilde{E}\{M\}, \tilde{u}\{M\}) \quad (13)$$

We accept, in this case, that the variable  $T_{\nu_{eff}} = \frac{M - \tilde{E}\{M\}}{\tilde{u}\{M\}}$  follows a Student distribution with the number of degrees of freedom to be determined.

To this end an interesting formula can be taken into account (Welch-Satterthwaite) [5], [6] and [7]

$$\nu_{eff} = \frac{\left[ \sum_{i=1}^r s_{oi}^2 + \sum_{j=1}^s u_j^2 \right]^2}{\sum_{i=1}^r \frac{s_{oi}^4}{n_i - 1} + \sum_{j=1}^s \frac{u_j^4}{\nu_j}} \quad (14)$$

where  $\nu_j$  is the number of degrees of freedom which can be attributed to  $u_j^2$ .

The contributions of the  $u_j^4$  to the denominator is sometimes negligible when its number of degrees of freedom is equal to infinite.

If the  $u_j$  were not acceptably known, we can use their estimates  $\tilde{u}_j$  with standard uncertainty  $\sigma(u_j)$ , to define following uncertainty interval (uncertainty of uncertainty):

$$\tilde{u}_j \left( 1 \pm \frac{\sigma(\tilde{u}_j)}{\tilde{u}_j} \right)$$

The ratio  $\frac{\sigma(\tilde{u}_j)}{\tilde{u}_j}$  can be made equal to the ratio  $\frac{\sigma(s_{oi}^-)}{s_{oi}^-}$  in the case of type A uncertainty components.

It's also possible refer to table E.1 of GUM [5] p.49, where the equivalent ratio  $\frac{\sigma(s(\bar{q}))}{\sigma(\bar{q})}$  is brought back as a percentage of the number of degree of freedom equal to  $(n-1)$ . There exists an approximate formula for this ratio when  $(n-1)$  is large enough given that  $\frac{\sigma(s(\bar{q}))}{\sigma(\bar{q})} \cong [2(n-1)]^{-1/2}$ .

By analogy with the above mentioned formula it is possible to obtain the number of degrees of freedom associated with  $u_j^2$ , equal to:

$$\nu_j = \frac{1}{2} \left[ \frac{\sigma(\tilde{u}_j)}{\tilde{u}_j} \right] \quad (15)$$

For example, if the relative standard uncertainty is equal to 25% in Eq. (15), we deduce  $\nu_j = 8$ .

Obviously, the smaller is the relative standard uncertainty the greater is the number of degrees of freedom which can be attributed to this last one.

The final level of uncertainty with a confidence level  $p$  is obtained in this case from the following equation

$$P\left\{ \tilde{E}\{M\} - t_p \tilde{u}_M \leq M \leq \tilde{E}\{M\} + t_p \tilde{u}_M \right\} = p \quad (16)$$

where for a stated confidence level, the corresponding value  $t_p$  is determined from the Student table corresponding to the number of degrees of freedom,  $\nu_{eff}$ , obtained from Eq. (14).

If  $\nu_{eff}$  is not integer, as is usually the case, then it is necessary to round it off to nearest lower integer.

### III. Conclusions

The aim of this paper is to introduce a new detailed method as a contribution for defining the Word Error Rate measurement and the Annex A of the IEEE Std 1241 [1]. This is made according with GUM [5] and its supplements carried out by BIPM. It has been quantified the uncertainty level with relative confidence level in the two cases of  $n$  successive observation and of the sum of  $n$  successive observation using Student and Chi-square distributions, with the final aim to give a contribution for a new draft of such standard.

### References

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